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## Bioelectrical impedance analysis to estimate one-repetition maximum measurement of muscle strength for leg press in healthy young adults

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Resistance training (RT) progress is determined by an individual's muscle strength, measured by one-repetition maximum (1RM). However, this evaluation is time-consuming and has some safety concerns. Bioelectrical impedance analysis (BIA) is a valid and easy-to-use method to assess skeletal muscle mass (SMM). Although BIA measurements are often correlated with muscle strength, few studies of 1RM for RT and BIA measurements are available. This observational study examined the relationship between 1RM and BIA measurements and developed BIA-based prediction models for 1RM. Thirty-five healthy young Japanese adults were included. SMM and the skeletal muscle mass index (SMI) were measured using the BIA device. In addition, dominant-leg 1RM was measured using a unilateral leg-press (LP) machine. The correlations between BIA measurements and 1RM were calculated, and simple regression analyses were performed to predict 1RM from the BIA variables. The results showed significant correlations between 1RM and dominant-leg SMM ( $R = 0.845$ ,  $P = 0.0001$ ) and SMI ( $R = 0.910$ ,  $P = 0.0001$ ). The prediction models for 1RM for LP derived from SMM of the dominant leg and SMI were  $Y = 8.21x + 8.77$  ( $P = 0.0001$ ),  $R^2 = 0.73$ , and  $Y = 15.53x - 36.33$  ( $P = 0.0001$ ),  $R^2 = 0.83$ , respectively. Our results indicated that BIA-based SMI might be used to predict 1RM for LP accurately.

Resistance training (RT) improves physical function and reduces the incidence of cardiovascular disease and all-cause mortality in adults<sup>1</sup>. In addition, multiple types of exercise, including RT, reportedly reduced fall risk in older adults<sup>2</sup>. Therefore, RT is an important physical activity to maintain health<sup>3</sup>. Particularly, RT progress is determined by one-repetition maximum (1RM), the maximum load a participant can lift at one repetition<sup>4</sup>. This measurement is useful for selecting strategies by either adopting heavier loads with fewer repetitions and sets<sup>5</sup> or adopting moderate loads with additional repetitions and sets<sup>6</sup>.

Despite its time-consuming nature, direct 1RM measurement is a reliable strength evaluation method<sup>7</sup>. However, safety considerations are necessary because joint injuries in older adults<sup>8</sup> and blood pressure elevations in young adults<sup>9</sup> have been reported with the use of loads close to 1RM during RT. Therefore, for RT prescription, the development of safe 1RM measurement is important for individuals of all age groups. In addition, estimations from several repetitions performed with submaximal loads<sup>10–12</sup> or maximal isometric muscle strength<sup>13,14</sup> were reported as indirect 1RM measurements. Although these methods could shorten the time required to

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|                          | All participants ( <i>n</i> = 35) | Men ( <i>n</i> = 18) | Women ( <i>n</i> = 17) | <i>P</i> -value |
|--------------------------|-----------------------------------|----------------------|------------------------|-----------------|
| Age, years               | 28.3 ± 3.8                        | 29.8 ± 3.1           | 26.7 ± 3.9             | 0.013           |
| Height, cm               | 163.3 ± 8.8                       | 170.3 ± 4.5          | 155.9 ± 5.4            | 0.0001          |
| Weight, kg               | 55.5 ± 8.7                        | 61.8 ± 6.5           | 48.7 ± 4.6             | 0.0001          |
| BMI, kg/m <sup>2</sup>   | 20.7 ± 2.1                        | 21.3 ± 2.0           | 20.1 ± 2.1             | 0.082           |
| Whole body FM, kg        | 10.9 ± 3.8                        | 10.0 ± 4.0           | 11.8 ± 3.6             | 0.176           |
| Dominant-arm SMM, kg     | 2.2 ± 0.6                         | 2.8 ± 0.3            | 1.6 ± 0.2              | 0.0001          |
| Non-dominant arm SMM, kg | 2.2 ± 0.6                         | 2.7 ± 0.3            | 1.6 ± 0.2              | 0.0001          |
| Dominant-leg SMM, kg     | 7.2 ± 1.6                         | 8.5 ± 0.9            | 5.7 ± 0.6              | 0.0001          |
| Non-dominant leg SMM, kg | 7.2 ± 1.6                         | 8.5 ± 0.9            | 5.7 ± 0.6              | 0.0001          |
| Trunk SMM, kg            | 19.5 ± 3.8                        | 22.8 ± 1.8           | 15.9 ± 1.4             | 0.0001          |
| SMI, kg/m <sup>2</sup>   | 6.9 ± 1.0                         | 7.7 ± 0.5            | 6.1 ± 0.4              | 0.0001          |
| 1RM for leg-press, kg    | 71.0 ± 16.5                       | 84.2 ± 8.4           | 57.1 ± 10.2            | 0.0001          |

**Table 1.** Participants' characteristics. Data are presented as the mean ± standard deviation. *1RM* one-repetition maximum, *BMI* body mass index, *FM* fat mass, *SMI* skeletal muscle mass index, *SMM* skeletal muscle mass.

measure maximal loads, there are some concerns about cardiovascular safety<sup>15</sup> because of the nearly maximal effort required for the estimation. Therefore, it is necessary to develop other safer methods for 1RM prediction.

Muscle strength is determined by a combination of morphological and neural factors<sup>16</sup>. Morphological factors such as skeletal muscle mass (SMM) and cross-sectional skeletal muscle area in the extremities have been well documented to have a moderate-to-strong correlation with muscle strength or power<sup>17–19</sup>, and evaluating these morphological factors might represent a viable option for muscle function assessment. Several methods can be used to accurately assess SMM, including magnetic resonance imaging, computed tomography, and dual-X-ray absorptiometry<sup>20</sup>. However, these modalities are not easily utilized because of their cost and environmental restrictions. Further, radiation exposure in computed tomography and dual-X-ray absorptiometry limits their use. In recent years, bioelectrical impedance analysis (BIA) has been widely used in both clinical and research settings. BIA is an easy-to-use, non-invasive, and relatively inexpensive method to assess body composition, including the SMM. The method uses alternating electrical current to measure the body's resistance at a designed frequency and can estimate the body composition based on the differences in tissue-specific electrical conductivity. The validity of BIA for SMM measurement has been confirmed, particularly among Asian populations in previous studies<sup>21–23</sup>. In this regard, the Asian Working Group for Sarcopenia recommends its use for sarcopenia diagnosis, a condition characterized by progressive loss of muscle mass and strength upon aging<sup>24</sup>.

Some studies have reported moderate-to-strong correlations between muscle strength and BIA parameters<sup>25–27</sup>. However, most of the strength measurements used in previous studies involved isometric conditions<sup>21,25–27</sup> instead of dynamic muscle strength (e.g., 1RM, which is essential for RT prescription). If the relationship between 1RM and parameters, such as SMM or skeletal mass index (SMI, an index of upper and lower extremity muscle mass adjusted by height), is confirmed and SMM can be accurately estimated using BIA parameters, a simpler and safer method to prescribe RT without any concerns could be developed. This study aimed to examine the relationship between the BIA measurements and 1RM for leg-press (LP), a typical multi-joint RT exercise of the lower extremities, and to develop BIA-based prediction models of 1RM for LP. Subgroup analysis was also conducted to reveal sex differences in the correlation and accuracy of the BIA-based prediction models of 1RM.

## Results

Forty participants were enrolled in this study. Some participants with a history of injury to the lower limbs (*n* = 1) or lumbar regions (*n* = 1) and those with low back pain (*n* = 3) were excluded. Ultimately, 35 participants (18 men and 17 women) were included in this study. The participants' characteristics are presented in Table 1. Between the two sexes, men had significantly more muscle mass and higher SMI than women. In addition, the 1RM for LP was significantly higher in men than in women (Table 1).

In all participants, the correlation coefficients between 1RM for LP and dominant-leg SMM and 1RM for LP and SMI as a BIA measurement were 0.845 (*P* = 0.0001) and 0.910 (*P* = 0.0001), respectively. In men, the correlation coefficients between 1RM and dominant-leg SMM and 1RM and SMI were 0.527 (*P* = 0.025) and 0.752 (*P* = 0.0001), respectively. Conversely, in women, the correlation coefficients between 1RM and dominant-leg SMM and 1RM and SMI were 0.310 (*P* = 0.225) and 0.613 (*P* = 0.009), respectively (Table 2).

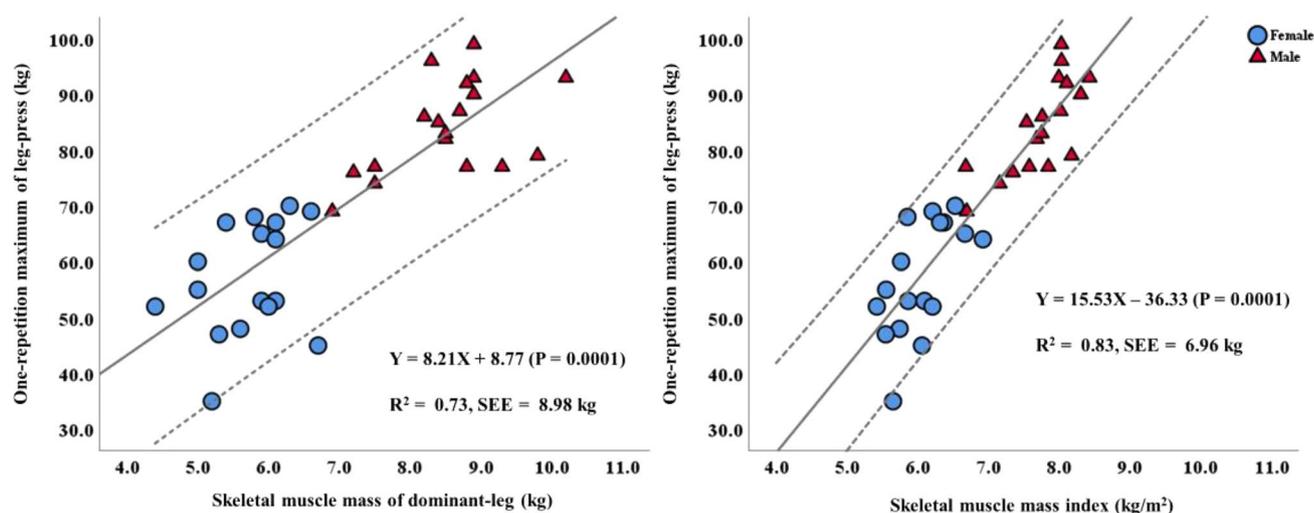
The results of the single linear regression analysis are presented in Table 3. Most of the BIA-based prediction models for 1RM for LP were statistically significant except for the model with dominant-leg SMM as an independent variable in women. The *R*<sup>2</sup> values of the prediction model using dominant-leg SMM and SMI in all participants were 0.73 (standard error of estimation [SEE]: 8.98 kg, *P* = 0.0001) and 0.83 (SEE: 6.96 kg, *P* = 0.0001), respectively (Fig. 1). In the 1RM prediction model that analyzed sex, *R*<sup>2</sup> values in men were 0.28 (SEE: 7.39 kg, *P* = 0.025) in dominant-leg SMM and 0.56 (SEE: 5.73 kg, *P* = 0.0001) in SMI. In women, the *R*<sup>2</sup> value of the prediction model with SMI as an independent variable was 0.38 (SEE: 8.29 kg, *P* = 0.0001).

|                  | 1RM for leg-press                 |                            |                           |
|------------------|-----------------------------------|----------------------------|---------------------------|
|                  | All participants ( <i>n</i> = 35) | Men ( <i>n</i> = 18)       | Women ( <i>n</i> = 17)    |
| Dominant-leg SMM | 0.845 ( <i>P</i> = 0.0001)        | 0.527 ( <i>P</i> = 0.025)  | 0.310 ( <i>P</i> = 0.225) |
| SMI              | 0.910 ( <i>P</i> = 0.0001)        | 0.752 ( <i>P</i> = 0.0001) | 0.613 ( <i>P</i> = 0.009) |

**Table 2.** Correlation analyses between BIA measurements and 1RM for leg-press. *1RM* one-repetition maximum, *BIA* bioelectrical impedance analysis, *SMI* skeletal muscle mass index, *SMM* skeletal muscle mass.

|                                   | Dependent variable | Prediction model     | R <sup>2</sup> | SEE  | 95% confidence interval |        | <i>P</i> -value |
|-----------------------------------|--------------------|----------------------|----------------|------|-------------------------|--------|-----------------|
|                                   |                    |                      |                |      | Lower                   | Upper  |                 |
| All participants ( <i>n</i> = 35) | Dominant-leg SMM   | $Y = 8.21X + 8.77$   | 0.73           | 8.98 | 6.798                   | 10.733 | 0.0001          |
|                                   | SMI                | $Y = 15.53X - 36.33$ | 0.83           | 6.96 | 13.022                  | 18.040 | 0.0001          |
| Men ( <i>n</i> = 18)              | Dominant-leg SMM   | $Y = 5.23X + 39.63$  | 0.28           | 7.39 | 0.763                   | 9.696  | 0.025           |
|                                   | SMI                | $Y = 12.74X - 14.25$ | 0.57           | 5.73 | 6.824                   | 18.647 | 0.0001          |
| Women ( <i>n</i> = 17)            | Dominant-leg SMM   | $Y = 5.15X + 27.56$  | 0.10           | 10.0 | -3.553                  | 13.831 | 0.225           |
|                                   | SMI                | $Y = 14.65X - 31.52$ | 0.38           | 8.29 | 4.265                   | 25.039 | 0.009           |

**Table 3.** Prediction models of 1RM for leg-press using BIA measurements. *1RM* one-repetition maximum, *BIA* bioelectrical impedance analysis, *SEE* standard error of estimation, *SMI* skeletal muscle mass index, *SMM* skeletal muscle mass.



**Figure 1.** Regression models for one-repetition maximum for leg-press from using BIA measurements. (a) Skeletal muscle mass of the dominant leg. (b) Skeletal muscle mass index. The dotted lines represent the 95% confidence interval.

## Discussion

This study aimed to determine the relationships between the 1RM for LP and BIA measurements and to develop a BIA-based prediction model of 1RM for LP in healthy young adults. We also examined sex differences in the correlational analysis and the prediction models' accuracy. There were two important findings in this study. First, there were strong correlations between the 1RM for LP and BIA measurements, and accurate prediction may be possible with BIA measurements, particularly with SMI as the dependent variable. Second, there were stronger correlations in men than in women; therefore, accurate 1RM estimation using BIA measurements might be attainable for men.

Although previous studies have revealed that isometric knee extensor muscle strength and BIA measurements are strongly correlated<sup>21,25–27</sup>, few studies analyzed the relationship between 1RM and dynamic muscle strength using BIA measurements. Kanada et al.<sup>28</sup> reported strong relationships ( $Rho = 0.70$  to  $0.78$ ,  $P < 0.01$ ) between the knee-extensor 1RM and each appendicular limb's SMM or whole-body SMM obtained using BIA in young men. In addition, these authors also reported that 1RM could be accurately estimated by combining isometric muscle strength and SMM in healthy young men<sup>28</sup>. Our results showed stronger correlation coefficients with 1RM for LP and BIA measurements than those in the study of Kanada et al., which might be related to the influence of

the required muscle activation in the adopted resistance exercise on BIA measurements. Knee extension exercise mainly requires quadriceps muscle activity<sup>29</sup>, whereas LP requires activities in most lower limb muscles<sup>30</sup>. Therefore, compared with the previous study, our method showed improved results because SMM or SMI, which includes most of the lower limb muscles, might be strongly correlated with and could accurately estimate the 1RM for the multi-joint LP exercise.

In our study, an accurate prediction model was developed when the SMI, an index of upper and lower extremity muscle mass adjusted by height, was used as a dependent variable rather than the dominant-leg SMM. This finding might be explained by the fact that all body parts, including the parts whose strength was directly measured as well as those enabling bodily stabilization at the time of measurement, affect strength assessment. Moreover, a previous study revealed that bodily stabilization by the hands influenced muscle strength production<sup>31</sup>. Since grasping handles is considered a body-stabilization strategy, which is not involved in physiological processes associated with remote muscle contractions<sup>32</sup>, the estimation accuracy of SMM and SMI as dependent variables to predict 1RM for LP might be influenced.

There were differences in the correlation coefficients and accuracy of the estimation model between men and women in this study. In this regard, the correlation between the 1RM for LP and BIA variables or the accuracy of the estimation model was higher in men than in women. A previous study reported a sex difference in the correlation between isometric knee muscle strength and BIA-based SMM in community-dwelling older adults<sup>33</sup>. Our study confirmed the sex differences between 1RM as dynamic muscle strength and BIA measurements and those relationships in healthy young adults. A possible explanation for these results may be the difference in neural determinants of muscle strength between the sexes. While it is well documented that men generally have greater muscle mass than women<sup>34,35</sup>, few studies of sex differences in neural factors are available. However, some studies have revealed that the firing rate of the vastus medialis differed between women and men<sup>36</sup>, and less steady force production in women was caused by unstable modulation of the motor firing discharge rate<sup>37</sup>. Moreover, BIA cannot evaluate such neural factors and only reflects morphological factors, which may explain the differences in the correlation and accuracy of the estimation model between the sexes in this study. Another possible explanation was that BIA could not evaluate the difference in the fiber characteristics of the lower limb muscles between men and women. A previous study revealed that type II fibers, which could produce more force than type I fibers, which are more suitable for continuous force production, are larger in men than in women<sup>38</sup>. Further, BIA could not distinguish muscle fiber types, which might have influenced the correlation and accuracy of the estimation model between the sexes in this study.

Our findings might help athletic trainers or fitness professionals resolve the concerns about the time-consuming and unsafe nature of 1RM measurements. Additionally, these results might potentially be applied in rehabilitation settings, where safety concerns are more important for future studies. Further, our findings might also indicate the necessity for practical equipment that could assess an individual's body characteristics or capacity more objectively to carry out physical training.

This study had some limitations. First, although sample size calculation was conducted prior to the study initiation, the sample size was too small to draw a clear conclusion regarding the validity and reproductivity of the established prediction models. Hence, a study with a larger sample size should be conducted to confirm our results and their validity and reproductivity in the future. Second, our target population was healthy young adults, not older adults or people with pain or past injuries. Since SMM might not play an important role in muscle strength among healthy older adults<sup>39</sup> and muscle strength was lower in people with a history of injuries or low back pain compared to those without those problems<sup>40,41</sup>, the correlations and estimation models in our study could not be directly adapted to these populations. To address this issue, multivariable estimation models, adaptive to any population, should be developed in future studies. Finally, since the absolute SMM value reportedly differed among body composition analyzers<sup>42</sup>, the relationship and accuracy of the estimation models may differ from those of other BIA equipment. Therefore, the correlation and accuracy of the estimation model using other equipment should be confirmed in future studies.

In conclusion, 1RM for LP and BIA measurements were strongly correlated, and accurate 1RM prediction from BIA measurements might be attainable in healthy young adults. This methodology might provide a new perspective for sports or fitness experts to resolve the safety and time-consuming concerns for 1RM measurements. The application of our results to rehabilitation medicine might also be expected in future studies.

## Methods

**Study design and ethical approval.** This cross-sectional study protocol was approved by the institutional ethics committee of Shinshu University (approval number: 3722). This study was conducted in accordance with the Declaration of Helsinki and was revised in 2013. All participants were informed of the study's aim, procedures, and potential risks and signed informed consent forms before their participation.

**Study participants.** Healthy adults working as medical staff at Kakeyu-Misayama Rehabilitation Center, Kakeyu Hospital, Japan, were conveniently recruited via a displayed poster between July 2017 and November 2017. The inclusion criteria were as follows: (1) age  $\geq 20$  and  $< 40$  years, (2) no history of injury to the spine or lower limbs, (3) no history of neurological diseases, (4) no pain at rest or during exercise, (5) no pregnancy or possible pregnancy, and (6) no cardiac pacemaker.

**Procedure.** First, the body composition was measured using BIA, and then the 1RM was measured. Both assessments were conducted at a fixed time on the same day. Participants were instructed to refrain from eating or drinking large amounts of water 4 h before the measurements and consuming alcohol 8 h before the measurement. Participants were also required not to undertake any intense exercise for 8 h before the measurements.

**BIA measurements.** BIA measurements were performed using a portable body composition analyzer Inbody 430 (Biospace, Korea), equipped with a terra-polar eight-point tactile electrode system. It uses three multi-frequencies (5 kHz, 50 kHz, and 250 kHz) to measure the impedance of the participant's appendicular muscles and trunk for the estimation of the body composition. The measurements by multi-frequencies are considered a better method for assessing muscle function than the single-frequency measurement<sup>26</sup>. A portable body composition analyzer from the Inbody models was confirmed as a reliable and valid tool to assess the SMM in healthy men and women and is considered to have sufficient ability to assess the body composition such as SMM, body fat, and body fat percent like other advanced models<sup>43</sup>. Moreover, the Inbody 430 has been widely used to assess SMM, especially for the Japanese population, in various studies, including large sample cohort studies<sup>44–48</sup>. After the participants wiped their soles off, they stood on the analyzer's platform, grasping the handles with both hands according to the manufacturer's guidance. The measurements took approximately 40 s to complete. The analyzer calculated the absolute muscle and fat mass, body fat percentage, and segmental muscle mass values (upper and lower limbs of both sides and trunk). We used dominant-leg SMM and SMI, which was the sum of the appendicular SMM obtained by dividing the participants' squared height ( $\text{kg}/\text{m}^2$ ), for the analyses because the SMI was reportedly correlated with muscle function in people with sarcopenia<sup>49</sup>.

**1RM measurement.** The 1RM measurement was performed using the participant's dominant leg with an LP resistance training machine (HUR, Finland). This resistance training machine allowed the participants to lift the loads unilaterally. The 1RM procedure was performed according to the American College of Sports Medicine guidelines<sup>50</sup>. All participants underwent a 5-min warm-up session using an ergo cycle bike before the measurements. The participants sat on the LP machine with their hip and knee joints fixed at approximately 90°, and the pelvis was stabilized by the belt. The participants were also required to hold handgrips placed on both side of the machine seat with each hand. The familiarization session with LP with light resistance for 8–10 repetitions was performed using perceived 50% 1RM. The measurements were started at loads of 80% 1RM. The load in the measurement was progressively changed by 3–10 kg until the participants could not lift the loads. The goal was to complete a maximal lift in five attempts, and 3–5 min of rest were provided between sets. All tests were performed by the same evaluator in the same order.

**Statistical analysis.** The sample size analysis was conducted using G\* Power software 3.1.9.4 (Heinrich Heine University, Dusseldorf, Germany). Since moderate-to-strong correlations between BIA measurements and isometric muscle strength of the lower limbs have been previously reported<sup>21,25–27</sup>, we set the alpha to 0.05, power to 0.8, and effect size to 0.5 and calculated the required minimum sample size to be  $n=26$ . The participants' characteristics were presented as the mean  $\pm$  standard deviation. First, we confirmed the normality of the obtained data using the Shapiro–Wilk test, and we also confirmed the homogeneity between the sexes by unpaired t-test. Second, we identified the correlations between each of the variables obtained from the body composition analyzer and the 1RM for LP by calculating Pearson's product-moment correlation coefficients because the normality of those measurements was confirmed. Finally, to create the 1RM prediction models, a simple linear regression analysis was performed with BIA measurements as independent variables. To evaluate the models' accuracy,  $R^2$  and SEE parameters were considered. All analyses were performed using SPSS version 25 (International Business Machine Corp., Armonk, NY, USA). Statistical significance was considered at  $P$ -values  $< 0.05$ .

## Data availability

The datasets used in this study are available from the corresponding author upon reasonable request.

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## Author contributions

K.M., S.K., and K.S. conceptualized this study. K.S., Y.K., M.I., M.M. and S.K. contributed to data collection. K.S. and K.M. contributed to the data analysis. K.S. and K.M. wrote the initial draft, and all authors contributed to the interpretation of the data and critically reviewed the manuscript. All authors approved the final version of the manuscript.

## Competing interests

The authors declare no competing interests.

## Additional information

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